

Detecting Earnings Management. An Analysis of Credit Institutions' (Banks) Trading in Hungary

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ABSTRACT: This paper examines evidence of Earnings Management (EM) in annual financial reports of foreign and domestic credit institutions' trading in Hungary, an ex-communist country and a Central European economy, by applying an alternative approach, the Distribution of the Ratios method. Analyses were performed with 14 banking specific ratios for the period of 1999-2015, by applying Burgstahler and Dichev (1997), Degeorge et al. (1991) models, Kolmogorov-Smirnov, Monte-Carlo, Benchmark and Distributional tests. Primary findings confirm that (a) the Distribution of the Ratios method demonstrates that apart from significant evidence of EM presence, managers do not always manipulate the same variables, or one area of the financial statements, but at their discretion, choose different areas to engage EM and in different periods, and (b) Credit Institutions engaged in EM in the period Prior to and After the 2008 financial crisis. Additional Benchmark Analysis results present weak evidence of EM and should be read with caution; however, Benchmark comparison should not be excluded from research.

KEYWORDS: Earnings Management, Credit Institutions, Distributional Test, Ratios

Introduction

After the collapse of the communist regime in 1989, Hungary went through a political and economic transition. By 1999, credit institutions' (banks) have been privatized, and in 2004, Hungary becomes full member of the European Union (EU). Earlier papers of Szapáry (2001) and Várhegyi (2008) reported on consolidation, privatization, but no evidence of a research that examined Banks annual accounts, specifically in search for any evidence of EM. This paper set to fill this gap by testing annual statements of foreign and domestic credit institutions' (banks) trading in Hungary by applying a new approach. The rest of the paper is designed as follows: Summary of Literature Review, Methodology, Hypotheses and Data, Results, Discussion and Conclusion.

Literature review

Over the past 30 years, researchers have been mainly investigating non-financial company's annual accounts that might have led to accounting manipulations, thus to EM. Financial industry has not been as thoroughly investigated as non-financials due to strict rules that govern the financial industry and high financial penalties for cooking the books, sadly, the managers' incentive to engage in EM is not lower, but higher, see e.g., Shen and Chih (2005, 2696-2697). Since the '80's, studies investigated EM by applying accrual and/or the standard discontinuity models, see e.g., Beneish (1999), Shen and Chih (2005), or Gore, Pope and Singh (2007). Lately, an attempt to increase the power of tests were by Baber, Kang and Li (2011) and Dechow, Hutton et al. (2012) papers who attempted to incorporate reversal accruals into their testing models. However, the authors could not increase the testing power of more than 40%, due to reversal accruals timing effect. One of the earliest studies which investigated Loan Loss Provisions (LLP), a specific accrual for financials, was by Beatty, Chamberlain and Magliolo (1995) who show weak evidence in favor of EM, whereas Beatty, Ke and Petroni (2002) present evidence of EM in publicly traded banks. From 2002 onwards, authors published papers showing evidence of EM within financials both at a cross-country and at a country level, see for example Beretka (2016, 12-98). Earlier studies examined EM by testing financials, mainly the LLP variable, and the non-financials data. These and earlier evidence suggest that accrual testing designs with few variables are weak in examining EM, see e.g., Dechow, Ge and Schrand's (2010) exploratory study that outlines the summary of strength and weaknesses of five accrual models linked to theories, which have been used to test evidence of EM. This paper proposes a new empirical approach to investigate EM by testing all available variables from the annual financial statements. This is discussed in the section below.

Hypotheses, Methodology and Data

Hypotheses

This paper test total of 14 variables to investigate evidence of EM of foreign and domestic credit institutions' (Banks) annual accounts trading in Hungary. Two hypotheses were tested:

H0_(a): Credit institutions' (Banks) in Hungary do not manage annual accounts.

H0_(b): Credit institutions' (Banks) in Hungary do not manage annual accounts 'Prior to' and/or 'After' the 2008 financial crisis.

Methodology

Tests were performed with Burgstahler and Dichev (1997), Degeorge, Patel and Zeckhauser (1999) models, Kolmogorov-Smirnov, Monte-Carlo Method, benchmark and distributional tests. A total of 14 banking specific variables are tested on an annual basis. The two testing approaches are, see e.g., Beretka (2016, 129), *the Distribution of Ratios Method*, it has two testing designs, the *EMI* and the *EM2* models that test Hypothesis **H0_(a)**: and **H0_(b)**:. This paper calculates the 14 ratios by applying the *EMI* model, or,

$$EM1 = (AO_i - EO_i) / SD_i \quad (1)$$

where, *EMI* is Earnings Management Model 1, and is equal to the Actual Observation (AO) in period (i) minus the Expected Observation (EO) in period (i). $SD_i = \text{Standard Deviation}$, or, $SD_i = [Np_i (1 - p_i) + \frac{1}{4} N (p_{i-1} + p_{i+1}) (1 - p_{i-1} - p_{i+1})]^{1/2}$, where, $SD_i = \text{Standard Deviation}$ of the difference in period (i); $p_i = \text{probability}$ of an observation in interval (i); $N = \text{number}$ of total sample; $Np_i = \text{total number}$ of estimated Standard Deviation (SD) in interval (i); $p_{i-1} = \text{number}$ in interval $i-1$; $p_{i+1} = \text{number}$ in interval $i+1$. For additional explanation see Beretka (2016, 300-304). By applying the *EM2* model, or,

$$EM2 = T_n = [\Delta p_n - \text{mean} (\Delta p_i)] / \text{s.d.} (\Delta p_i) \quad (2)$$

where, $i \in R$, $i \neq n$. p_i is the ratio of the actual sample for year i of banks years, Δp_n is the difference of $p_i - p_{i-1}$. $\text{Mean} (\Delta p_i)$ is the average of Δp but excluding p_i and $\text{s.d.} (\Delta p_i)$ is the standard deviation of Δp , excluding Δp_i .

Test statistics with Kolmogorov-Smirnov (K-S) were performed, see, e.g., Massey (1951) for the Kolmogorov-Smirnov (K-S) test statistics, as well Lilliefors (1967, 399). K-S is applicable for ratio or interval data, a sample of N observations: $D = \max x \in R | F(x) - F_0(x) |$, where, $F(x)$ is the cumulative normal distribution, and $F_0(x)$ is the sample cumulative distribution, with $\mu = \bar{x}$ sample mean and $\sigma^2 = s^2$ sample variance with denominator $n-1$. $F(x) = F_0(x)$ for all x from $-\infty$ to $+\infty$; $F(x) \neq F_0(x)$ for at least one value of (x). The K-S one sample test is non-parametric and distribution-free, an exact test, see, e.g., Panik (2005, 570). Additionally, benchmark analysis was calculated for each 14 ratio. Benchmark formula reads as:

$$\bar{x}_{i,j} = \sum x_{i,j} / N_{i,j} \quad (3)$$

where, $\bar{x}_{i,j} = \text{Is the Mean}$; $\sum x_{i,j} = \text{The sum of all 'x' values}$; $N_{i,j} = \text{The number of 'x' values}$, $i, j = \text{Ratios and years respectively}$. Papers on benchmarking of non-financials are Kent and Routledge (2015); Sun and Rath (2012); Dechow, Ge and Schrand (2010); Crump and Teeguarden (2009);

Habib (2007); Dattakumar and Jagadeesh (2003); Yasin (2002). Testing sample on an annual basis with the *EMI* and *EM2* models for 'Prior to' and 'After' the 2008 financial crisis to test **H0_(b)**: with K-S test statistics with 14 banking specific ratios. The tested 14 ratios for **H0_(a,b)**: are:

- *Debt to Equity (DTE)* = Total Liabilities_(t) divided by Equity_(t).
 - *Equity to Loans (ETL)* = Average Equity_(t) divided by Loans_(t). Where: Average Equity_(t) = (Equity_(t) + Equity_(t-1)) / 2
 - *Loans to Deposits Ratio (LTD)* = Loans_(t) divided by Deposits_(t).
 - *Loans to Total Assets (LTA)* = Loans_(t) divided by Total Assets_(t).
 - *Gross Yield on Earning Assets (GYEA)* = Tot. Interest Income_(t) divided by Tot. Earning Assets_(t).
 - *Rate Paid on Funds (RPF)* = Total Interest Expenses_(t) divided by Total Earning Assets_(t).
 - *Sales Growth Index (SGI)* = Sales_(t) divided by Sales_(t-1).
 - *Interest (Sales) Receivables Index (IRI)* = [Accrued Interest receivables divided by sales_(t)] divided by [Accrued Interest receivables divided by sales in receivables_(t-1)].
 - *Gross Margin Index (GMI)* = Gross Margin_(t-1) divided by Gross Margin_(t). Where: Gross Margin_(t) = (Total interest income_(t) - Tot Int. expenses_(t)) divided by Total Interest Income_(t).
 - *Net interest Margin (NIM)* = Net Interest Income_(t) divided by Average Earning Assets_(t-1). Where: Net interest income = Total Interest Income_(t) - Total Interest Expenses_(t). Average Earning Assets correspond to all assets that earn income, assets.
 - *Profit Margin (PATM)* = Profit After Tax (PAT)_(t) divided by Net Interest (sales)_(t).
 - *Return on Equity (ROE)* = Profit After Tax_(t) divided by Average Equity_(t).
 - *Return on Asset (ROA)* = Profit After Tax_(t) divided by Average Total Assets_(t).
- Where: Average Total Assets = (Total Assets_(t) + Total Assets_(t-1)) divided by 2
- *Equity to Total Assets (EtA)* = Equity_(t) divided by Average Total Assets_(t).
- See, e.g., Beretka (2016, 142) for ratio calculations. Ratios are calculated for the same companies within a time frame, for example, company ‘z’ in year (t) is calculated with the same company ‘z’ in year (t-1) or in (t-2).

Data

Audited data of publicly and privately owned foreign and domestic Credit Institutions – banks (operating as Joint-Stock Companies in Hungary) were obtained from the Central Bank of Hungary – CBH and from the Hungarian Financial Supervisory Authority (HFSA). Since 1 October 2013, HFSA has become part of the CBH and operates under the umbrella of CBH, CBH/HFSA thereafter. Tested data is from 1999 to 2015 and consist of one listed (on the Budapest Stock Exchange; OTP Bank) and non-listed banks. For the period 1999-2015, CBH/HFSA published data annually and per Hungarian Accounting Standards (HAS); including for the listed OTP Bank. See HAS vs. IFRS differences in Table 3.1, Beretka (2016, 113-116). From 2017, CBH/HFSA introduces IFRS reporting, however, as per 2017 published data by the CBH/HFSA, only few banks switched from HAS to IFRS. For research purposes, CBH/HFSA gather individual companies (audited) annual accounts and create their own simplified version of the annual reports for each financial company, shown in Hungarian Forint (HUF). Hungary is not member of the European Monetary Union. Tested annual data consists of Profit and Loss Account and Balance Sheet. Cash Flow is not prepared nor published by CBH/HFSA. Number of individual banks per year is between 37 to 45 and per variable varies from 502 to 682 for the sample. This paper performed tests of Credit Institutions – banks annual accounts that were prepared under the HAS.

Empirical Results

Statistical results of Ratios tested on the Annual basis with the EM1 model

The Distribution of Ratio method is a hands on approach, based on statistical and distributional

Table 1. One-Sample Kolmogorov-Smirnov (K-S) test run for all ratios on an annual basis with the *EMI* model with $p = 0.01$ Significance level

Equity to Total Assets (ETA)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	29	29	29	29	28	28	26	26	26	31	27	27	27	30
Mean	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00	-.02	.05	-.02
Std. Deviation	.01	.01	.00	.01	.01	.01	.00	.00	.00	.00	.00	.08	.23	.07
Kolmogorov-Smirnov Z	1.36	.93	1.11	1.22	1.81	1.56	1.28	.93	1.10	.99	.57	2.45	2.55	2.46
Asymp. Sig. (2-tailed)	.050	.357	.171	.104	.003	.016	.075	.349	.174	.284	.899	.000	.000	.000
Monte Carlo Sig. (2-tailed)	.036	.321	.163	.089	.000	.005	.064	.314	.168	.268	.867	.000	.000	.000
Equity to Loans (ETL)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	31	26	30	29	29	28	28	28	28	31	29	29	29	32
Mean	.00	.00	.01	-.01	.03	.00	.00	.00	.00	.00	.00	-.02	.06	-.03
Std. Deviation	.02	.06	.09	.15	.17	.05	.11	.05	.02	.01	.01	.12	.35	.12
Kolmogorov-Smirnov Z	1.07	1.34	2.13	2.46	2.60	2.05	2.24	1.97	1.52	1.15	.70	2.73	2.75	2.23
Asymp. Sig. (2-tailed)	.199	.055	.000	.000	.000	.000	.000	.001	.020	.141	.705	.000	.000	.000
Monte Carlo Sig. (2-tailed)	.165	.029	.000	.000	.000	.000	.000	.000	.010	.118	.666	.000	.000	.000
Gross Margin Index (GMI)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	39	39	36	36	36	35	35	35	36	35	30	30	30	35
Mean	-.02	-.01	.01	.02	-.03	-.01	.01	.01	.01	-.04	.05	-.01	.02	-.01
Std. Deviation	.03	.03	.04	.06	.05	.05	.06	.07	.06	.06	.15	.14	.12	.14
Kolmogorov-Smirnov Z	.56	.46	.73	1.35	.51	.71	1.35	.92	1.30	1.30	1.52	1.00	.83	1.18
Asymp. Sig. (2-tailed)	.910	.982	.659	.053	.956	.694	.053	.359	.068	.069	.019	.274	.501	.123
Monte Carlo Sig. (2-tailed)	.848	.965	.598	.045	.924	.610	.033	.310	.057	.047	.025	.238	.446	.109
Gross Yield on Earning Assets (GYEA)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	43	43	41	40	40	40	40	39	37	37	37	37	37	37
Mean	.00	.00	.02	-.01	.02	.01	-.03	.07	-.02	-.01	.02	.04	-.04	.00
Std. Deviation	.05	.05	.05	.07	.13	.10	.13	.39	.14	.05	.06	.06	.03	.01
Kolmogorov-Smirnov Z	1.44	1.32	1.34	.96	1.79	1.43	1.78	2.52	1.92	.72	1.08	.86	1.25	1.35
Asymp. Sig. (2-tailed)	.032	.060	.056	.310	.003	.034	.004	.000	.001	.670	.198	.452	.088	.052
Monte Carlo Sig. (2-tailed)	.027	.049	.040	.268	.002	.027	.002	.000	.000	.639	.175	.394	.077	.046

analysis of all computable ratios from the financial statements. *Table 1* presents K-S test statistics with Monte with Monte Carlo simulation for each 14 ratios with 99% Confidence Interval levels, as well as with $p = 0.01$ Significance levels. *Table 1.* presents test outputs of each ratio for the period from 1999 to 2015, with low Mean and S.D. outputs, and negative and positive values. Test results for

Table 1. Cont.

Interest (Sales) Receivables Index (IRI)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	35	35	35	35	35	35	34	34	34	35	30	30	30	35
Mean	.03	.05	-.04	.03	.02	-.04	.04	-.01	.01	-.01	.07	-.12	.04	.00
Std. Deviation	.30	.43	.14	.13	.17	.09	.13	.11	.09	.12	.11	.08	.09	.15
Kolmogorov-Smirnov Z	1.95	2.21	1.69	1.00	1.73	1.20	.94	1.28	.53	.78	1.18	1.18	.50	1.06
Asymp. Sig. (2-tailed)	.001	.000	.007	.268	.005	.113	.341	.076	.944	.580	.125	.124	.961	.208
Monte Carlo Sig. (2-tailed)	.000	.000	.000	.239	.000	.091	.297	.053	.922	.530	.100	.097	.953	.174
Loans to Total Assets (LTA)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	38	34	41	36	34	34	35	36	35	35	37	38	37	37
Mean	.02	-.04	.02	.00	.00	.00	.01	.01	-.01	.00	.00	.01	.00	.00
Std. Deviation	.04	.04	.03	.02	.03	.05	.06	.06	.05	.05	.05	.05	.05	.05
Kolmogorov-Smirnov Z	.53	.60	.53	1.17	.81	1.17	.98	.39	.65	.54	.78	.59	.96	.79
Asymp. Sig. (2-tailed)	.939	.869	.941	.130	.523	.131	.288	.998	.799	.931	.585	.883	.321	.557
Monte Carlo Sig. (2-tailed)	.897	.813	.899	.097	.450	.103	.264	1.000	.744	.886	.503	.824	.278	.481
Loans to Deposits Ratio (LTD)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	24	16	23	28	29	29	29	2	2	30	31	31	31	31
Mean	1.39	.68	-.09	.04	.04	.01	.05	.04	.09	.01	.04	.00	-.11	.06
Std. Deviation	7.14	1.29	.41	.39	.38	.44	.57	.51	.53	.26	.28	.13	.68	.38
Kolmogorov-Smirnov Z	2.20	1.26	0.97	1.63	1.38	0.99	0.87	1.13	1.35	1.40	1.32	0.80	2.30	1.79
Asymp. Sig. (2-tailed)	.000	.084	.305	.010	.043	.277	.436	.153	.053	.041	.061	.552	.000	.003
Monte Carlo Sig. (2-tailed)	.000	.057	.244	.008	.028	.234	.362	.129	.044	.018	.054	.519	.000	.003
Net interest Margin (NIM)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	38	38	36	36	36	35	35	35	35	35	30	30	30	35
Mean	.00	.00	.00	.02	-.01	.00	.00	.01	-.01	.00	.00	.03	-.03	.00
Std. Deviation	.05	.04	.05	.10	.05	.04	.06	.06	.04	.04	.07	.06	.03	.02
Kolmogorov-Smirnov Z	1.35	.97	.97	1.59	1.28	.88	1.30	.89	.96	1.40	.99	.72	.83	1.41
Asymp. Sig. (2-tailed)	.052	.298	.307	.013	.074	.421	.067	.409	.316	.040	.286	.683	.493	.038
Monte Carlo Sig. (2-tailed)	.037	.244	.267	.008	.076	.368	.066	.357	.277	.045	.240	.599	.450	.039

Asymp. Sig. and Monte Carlo Sig. statistically significant years for ETA are: 2001-2003 and 2010; for ETL: 2001-2003 and 2007-2012; for GMI, the results are statistically insignificant; for GYEA are: 2006-2008 and 2010; for IRI: 2010 and 2012-2014; for LTA, the results are statistically insignificant; for LTD are: 2001-2001, 2011, 2014; for NIM: 2011 for Monte Carlo Sig. only; for PATM: 2004-2005 and 2007; for ROA are:

Table 1. Cont.

Profit Margin (PATM)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	39	41	41	38	38	38	38	40	36	37	34	34	34	33
Mean	-.13	.12	.00	-.01	.01	.03	-.01	.00	-.02	-.01	.08	.03	-.01	.04
Std. Deviation	.44	.39	.18	.12	.17	.19	.17	.27	.08	.18	.43	.17	.29	.47
Kolmogorov-Smirnov Z	1.47	1.51	1.04	.79	1.00	1.44	1.38	1.99	1.16	2.00	2.59	1.32	.75	1.14
Asymp. Sig. (2-tailed)	.027	.021	.227	.561	.273	.032	.044	.001	.135	.001	.000	.063	.628	.148
Monte Carlo Sig. (2-tailed)	.015	.015	.194	.488	.246	.013	.025	.000	.121	.000	.000	.038	.580	.109
Return on Asset (ROA)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	38	38	37	37	37	35	35	35	37	35	31	31	31	35
Mean	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00
Std. Deviation	.01	.01	.01	.01	.01	.01	.00	.01	.01	.00	.00	.00	.01	.00
Kolmogorov-Smirnov Z	1.02	1.10	.87	.78	1.03	1.06	1.05	1.76	1.83	1.03	.79	1.04	1.87	1.69
Asymp. Sig. (2-tailed)	.253	.175	.438	.583	.240	.215	.218	.004	.003	.239	.568	.234	.002	.007
Monte Carlo Sig. (2-tailed)	.220	.134	.396	.535	.191	.185	.187	.002	.000	.217	.500	.195	.000	.002
Return on Equity (ROE)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	21	15	18	18	16	18	17	18	17	20	21	20	21	25
Mean	-.17	-.14	-.16	-2.82	-1.90	-2.13	.54	-2.55	-.20	.11	-.18	.33	-.29	-.34
Std. Deviation	8.67	1.58	3.57	5.18	8.38	6.52	3.66	10.70	.72	.86	1.27	.92	.78	3.69
Kolmogorov-Smirnov Z	1.53	.76	.87	1.40	1.31	1.48	1.11	1.73	.96	1.36	1.07	1.11	1.28	1.76
Asymp. Sig. (2-tailed)	.018	.615	.431	.041	.065	.025	.171	.005	.317	.049	.203	.166	.074	.004
Monte Carlo Sig. (2-tailed)	.011	.558	.381	.038	.068	.023	.158	.008	.294	.049	.230	.196	.068	.004
Rate Paid on Funds (RPF)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	37	38	37	37	37	38	38	38	37	37	37	35	35	35
Mean	.00	.00	.00	.00	.00	.00	.00	-.02	.11	-.03	.00	.00	.00	.00
Std. Deviation	.00	.00	.00	.00	.00	.01	.02	.18	.72	.18	.01	.01	.00	.00
Kolmogorov-Smirnov Z	.76	.95	.81	.85	.97	1.35	1.90	2.99	3.23	3.05	1.07	1.17	1.04	1.25
Asymp. Sig. (2-tailed)	.617	.328	.534	.460	.306	.051	.001	.000	.000	.000	.202	.132	.228	.089
Monte Carlo Sig. (2-tailed)	.552	.300	.481	.413	.279	.052	.000	.000	.000	.000	.182	.107	.194	.074

2001-2002 and 2007; for ROE: 2001 and 2007; for RPF: 2005-2008; for SGI are: 2007-2009 and 2001 for Monte Carlo Sig only.

Table 1. Cont.

Sales Growth Index (SGI)	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001
N	38	38	36	36	36	35	35	35	35	37	30	30	30	35
Mean	.01	-.02	.01	.01	-.04	.04	.00	-.06	.22	-.13	-.66	2.18	-.58	-.01
Std. Deviation	.07	.04	.06	.06	.09	.20	.23	.32	.20	.09	.39	1.59	.58	.30
Kolmogorov-Smirnov Z	1.25	.79	.99	1.51	1.33	2.07	1.92	2.11	1.19	.89	.74	1.33	1.13	1.53
Asymp. Sig. (2-tailed)	.088	.559	.285	.021	.058	.000	.001	.000	.117	.405	.642	.058	.158	.018
Monte Carlo Sig. (2-tailed)	.045	.502	.265	.004	.031	.000	.000	.000	.097	.356	.591	.045	.105	.008

Finance managers' skills have become so sophisticated in manipulating accounting entries that evidence of EM may not emerge only in one section of the financial statements, but evidence may arise in multiple, or in all areas of financial statements and in diverse time frames. *Table 1* presents this evidence, and as it was expected, not every year shows statistically significant test results, which suggests that EM is not present in every financial year for each ratio. This is in line with practise, as managers in order to hide they action, as well to avoid high financial penalties, they will engage in EM in different periods. The same is true for 95% Confidence Interval and at $p = 0.05$ Significance levels for the *EMI* model. Additional test statistics on annual basis for 95% Confidence Interval and at $p = 0.05$ Significance levels for the *EMI* model are available from the author.

Table 2. One-Sample Kolmogorov-Smirnov test run on an Annual basis with the *EM2* model, with $p = 0.01$ Significance level.

Debt to Equity (DTE)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	2000
N	32	33	32	32	30	30	30	30	30	31	31	33	33	32	32	34
Mean	-.01	.04	.00	-.02	-.12	.09	-.55	.37	.10	.07	.03	-.05	-.01	.04	.03	-.01
Std. Deviation	.30	.51	.37	.04	.64	.71	2.98	2.05	1.32	.53	.13	.32	.29	.23	.22	.05
Kolmogorov-Smirnov Z	1.53	1.78	1.80	1.83	2.46	2.11	2.50	2.82	2.20	2.34	1.78	2.27	1.70	2.04	1.87	1.75
Asymp. Sig. (2-tailed)	.019	.003	.003	.003	.000	.000	.000	.000	.000	.000	.003	.000	.006	.000	.002	.004
Monte Carlo Sig. (2-tailed)	.026	.002	.008	.004	.000	.000	.000	.000	.000	.000	.004	.000	.006	.000	.004	.008
Equity to Total Assets (ETA)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	
N	31	29	29	29	29	28	28	26	26	31	30	27	27	30	32	
Mean	.00	-.01	.01	.00	-.02	.02	.00	.00	.00	.00	-.01	.00	-.57	.51	.02	
Std. Deviation	.08	.09	.07	.07	.15	.15	.06	.03	.06	.05	.02	.06	2.81	2.66	.17	
Kolmogorov-Smirnov Z	1.39	1.31	1.06	1.11	1.74	1.92	1.70	1.40	1.10	1.17	1.34	0.50	2.55	2.53	1.80	
Asymp. Sig. (2-tailed)	.042	.064	.207	.172	.005	.001	.006	.040	.177	.127	.055	.967	.000	.000	.003	
Monte Carlo Sig. (2-tailed)	.023	.065	.190	.157	.007	.000	.009	.021	.146	.095	.056	.944	.000	.000	.000	

Table 2. Cont.

Equity to Loans (ETL)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	
N	31	29	29	30	29	29	28	28	28	31	31	29	29	32	32	
Mean	.63	-.05	-.05	-.05	-.06	-.04	-.04	-.04	-.05	-.05	-.05	-.05	-.07	-.03	-.05	
Std. Deviation	3.78	.01	.03	.04	.10	.04	.07	.03	.01	.01	.00	.00	.10	.10	.03	
Kolmogorov-Smirnov Z	2.99	1.48	2.08	2.13	2.64	2.67	2.42	2.44	1.62	1.56	1.13	0.93	2.76	2.78	2.05	
Asymp. Sig. (2-tailed)	.000	.025	.000	.000	.000	.000	.000	.000	.010	.015	.155	.347	.000	.000	.000	
Monte Carlo Sig. (2-tailed)	.000	.025	.000	.000	.000	.000	.000	.000	.004	.007	.124	.306	.000	.000	.000	
Gross Margin Index (GMI)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	
N	39	39	39	36	36	37	35	35	36	37	35	30	30	35	38	
Mean	.22	-.12	-.22	-.22	.41	-.15	-.39	-.02	.08	.24	-.64	.01	.28	.40	.14	
Std. Deviation	.33	.34	.42	.54	.70	.48	.62	.63	.77	.54	1.21	1.85	1.57	1.48	1.54	
Kolmogorov-Smirnov Z	0.87	0.92	1.07	1.20	1.15	0.63	1.15	0.62	1.24	1.44	1.51	1.28	1.28	1.23	1.50	
Asymp. Sig. (2-tailed)	.431	.365	.200	.115	.143	.825	.145	.831	.092	.031	.021	.075	.077	.097	.022	
Monte Carlo Sig. (2-tailed)	.358	.292	.171	.095	.117	.782	.114	.782	.074	.019	.011	.063	.063	.071	.013	
Gross Yield on Earning Assets (GYEA)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	2000
N	40	43	41	40	40	40	40	40	39	37	37	38	37	37	40	42
Mean	-.03	-.10	-.20	.03	-.17	.03	.11	-.46	.38	-.08	-.16	.10	.65	.01	-.02	-.04
Std. Deviation	.39	.37	.47	.45	.90	1.02	.64	2.37	2.42	.57	.47	.53	.44	.12	.05	.08
Kolmogorov-Smirnov Z	1.11	1.39	1.39	1.05	1.55	1.65	1.43	2.60	2.37	1.05	1.09	0.90	1.11	1.61	1.19	1.40
Asymp. Sig. (2-tailed)	.171	.042	.041	.224	.016	.009	.033	.000	.000	.225	.183	.389	.173	.011	.119	.039
Monte Carlo Sig. (2-tailed)	.119	.027	.017	.165	.010	.006	.010	.000	.000	.193	.162	.350	.152	.003	.070	.022
Interest (Sales) Receivables Index (IRI)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	
N	38	37	37	35	35	36	35	35	35	37	35	29	29	35	38	
Mean	-.04	.01	.07	-.06	.02	.06	-.07	-.44	.44	.05	.01	.15	-.13	-.03	-.04	
Std. Deviation	.29	.18	.41	.14	.25	.19	.14	2.77	2.58	.22	.15	.13	.13	.12	.32	
Kolmogorov-Smirnov Z	1.92	1.13	2.22	1.21	1.41	1.35	1.19	2.77	2.74	1.10	1.06	1.12	0.89	0.72	1.74	
Asymp. Sig. (2-tailed)	.001	.159	.000	.106	.037	.053	.116	.000	.000	.179	.215	.161	.407	.675	.005	
Monte Carlo Sig. (2-tailed)	.002	.137	.000	.101	.040	.042	.112	.000	.000	.152	.205	.131	.350	.656	.006	

Table 2. Cont.

Loans to Total Assets (LTA)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	2000
N	38	33	33	38	37	37	37	38	38	37	37	38	37	37	39	42
Mean	.16	.34	-.59	-.26	-.07	-.13	-.09	.05	.16	-.15	.03	-.08	.02	.25	.15	.17
Std. Deviation	.92	.85	1.03	1.00	.69	.95	1.09	1.06	1.15	1.31	1.13	1.00	.83	1.07	.94	.62
Kolmogorov-Smirnov Z	0.50	0.59	0.77	1.10	1.20	1.18	0.96	0.89	0.74	0.55	0.76	0.56	0.67	0.79	0.79	1.23
Asymp. Sig. (2-tailed)	.962	.877	.588	.176	.111	.125	.314	.411	.647	.922	.606	.911	.766	.554	.554	.095
Monte Carlo Sig. (2-tailed)	.941	.834	.552	.133	.092	.099	.265	.336	.586	.899	.532	.876	.711	.487	.512	.060
Loans to Deposits (LTD)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	2000
N	32	31	31	30	30	30	30	29	29	31	31	31	31	31	34	35
Mean	.00	-.66	.65	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00	.00
Std. Deviation	.06	2.79	2.80	.03	.02	.03	.01	.02	.02	.01	.01	.01	.00	.01	.01	.01
Kolmogorov-Smirnov Z	1.94	2.66	2.66	2.10	1.42	1.87	1.03	.88	1.26	1.64	1.53	.98	1.33	1.71	1.69	2.24
Asymp. Sig. (2-tailed)	.001	.000	.000	.000	.034	.002	.234	.418	.080	.009	.018	.290	.056	.005	.007	.000
Monte Carlo Sig. (2-tailed)	.000	.000	.000	.000	.028	.002	.234	.405	.056	.002	.022	.280	.065	.002	.008	.000
Net interest Margin (NIM)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	
N	39	38	38	36	36	37	35	35	36	37	35	30	30	35	38	
Mean	-.06	-.09	-.04	-.37	.31	-.06	-.19	-.11	.17	-.14	-.08	-.22	1.02	.02	-.01	
Std. Deviation	.83	.90	.49	1.55	1.65	.29	1.05	1.11	1.11	.30	1.17	1.28	.77	.30	.24	
Kolmogorov-Smirnov Z	1.80	1.44	0.74	1.48	1.52	0.65	1.14	1.09	0.86	1.63	1.17	0.76	0.93	1.49	1.70	
Asymp. Sig. (2-tailed)	.003	.031	.636	.024	.019	.800	.149	.184	.458	.010	.129	.605	.348	.024	.006	
Monte Carlo Sig. (2-tailed)	.000	.030	.568	.024	.019	.763	.121	.155	.436	.009	.105	.589	.342	.024	.004	
Profit Margin (PATM)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	2000
N	40	43	41	41	38	38	39	40	40	37	37	34	34	34	39	41
Mean	.15	-.32	.03	-.03	-.10	-.13	-.07	-.05	-.15	-.14	-.24	.19	.34	.00	-.05	.61
Std. Deviation	1.10	.96	.74	.33	.50	.49	.60	.85	.87	.40	1.39	1.17	.61	1.22	1.54	1.66
Kolmogorov-Smirnov Z	1.11	0.94	0.92	1.42	0.72	1.13	1.40	1.81	2.22	1.32	2.53	2.23	0.91	0.75	0.81	1.59
Asymp. Sig. (2-tailed)	.171	.340	.365	.035	.679	.156	.039	.003	.000	.060	.000	.000	.380	.634	.532	.013
Monte Carlo Sig. (2-tailed)	.177	.310	.362	.023	.628	.141	.026	.000	.000	.047	.000	.000	.333	.593	.494	.005

Table 2. Cont.

Return on Asset (ROA)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	
N	39	38	38	37	37	37	35	35	37	37	35	31	31	35	37	
Mean	.38	-.48	.30	.00	-.01	-.24	-.05	.09	-.24	-.09	.04	.02	.38	.00	-.05	
Std. Deviation	1.07	1.50	1.25	1.36	1.12	.81	.70	.79	1.44	.23	.62	.61	.92	.82	.33	
Kolmogorov-Smirnov Z	1.23	1.11	1.26	0.92	0.94	1.45	1.26	1.32	1.90	1.27	0.77	0.76	1.69	1.85	1.52	
Asymp. Sig. (2-tailed)	.097	.173	.085	.359	.344	.030	.083	.061	.002	.080	.598	.609	.007	.002	.020	
Monte Carlo Sig. (2-tailed)	.067	.173	.085	.317	.297	.020	.091	.061	.000	.067	.570	.596	.002	.000	.019	
Return on Equity (ROE)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	
N	31	29	29	32	30	28	27	27	29	32	31	29	29	32	33	
Mean	.09	-.08	-.02	.03	.05	-.02	-.10	.17	-.24	.04	.01	.04	.03	.03	-.03	
Std. Deviation	.44	.35	.19	.48	1.06	1.07	.32	2.57	2.52	.43	.13	.20	.16	.18	.24	
Kolmogorov-Smirnov Z	2.01	1.65	1.24	1.77	2.01	1.91	1.86	2.36	2.13	2.29	1.48	1.86	1.53	2.33	2.57	
Asymp. Sig. (2-tailed)	.001	.009	.092	.004	.001	.001	.002	.000	.000	.000	.025	.002	.018	.000	.000	
Monte Carlo Sig. (2-tailed)	.000	.007	.076	.002	.000	.000	.002	.000	.000	.000	.018	.002	.013	.000	.000	
Rate Paid on Funds (RPF)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	2000
N	37	39	38	37	37	37	38	40	38	37	37	38	35	35	39	41
Mean	.00	-.01	-.01	.00	.00	-.02	-.02	-.01	-.42	.46	-.01	.01	.04	.00	.00	.00
Std. Deviation	.02	.02	.03	.03	.02	.04	.05	.24	2.82	2.84	.03	.04	.03	.01	.00	.01
Kolmogorov-Smirnov Z	0.85	0.68	0.59	0.93	0.62	1.47	0.82	2.72	3.22	3.23	0.90	1.03	1.10	1.62	0.98	1.51
Asymp. Sig. (2-tailed)	.461	.747	.883	.356	.840	.027	.519	.000	.000	.000	.392	.236	.180	.010	.295	.021
Monte Carlo Sig. (2-tailed)	.388	.695	.833	.292	.781	.022	.469	.000	.000	.000	.328	.207	.181	.013	.272	.020
Sales Growth Index (SGI)	2015	2014	2013	2012	2011	2010	2009	2008	2007	2006	2005	2004	2003	2002	2001	
N	40	38	38	36	36	37	36	35	35	37	37	30	30	35	38	
Mean	.00	.00	-.02	-.02	.04	-.09	-.01	-.04	-.20	.30	-.06	-2.62	2.45	.16	.07	
Std. Deviation	.10	.10	.09	.09	.08	.23	.25	.50	.42	.26	.11	1.29	1.43	.56	.29	
Kolmogorov-Smirnov Z	1.35	1.08	1.15	1.32	1.47	2.10	1.59	2.40	1.61	1.16	1.07	0.68	0.83	1.86	1.60	
Asymp. Sig. (2-tailed)	.052	.196	.145	.060	.026	.000	.013	.000	.011	.134	.202	.748	.498	.002	.012	
Monte Carlo Sig. (2-tailed)	.030	.154	.110	.041	.015	.000	.004	.000	.002	.095	.160	.706	.437	.000	.004	

Notes: Statistically significant years with 99% Confidence Interval and with $p = 0.01$ Significance level for DTE are: 2000-2014; for ETA: 2001-2003, 2009-2011; for ETL: 2001-20013, 2007-2013, 2015, and additionally 2006 for Monte Carlo Sig. only; for GMI are statistically insignificant; for GYEA significant years are: 2007-2008, 2010, and 2002, 2009, 2011 for Monet Carlo Sig. only; for IRI: 2001, 2007-2008, 2013, 2015; for LTA are statistically insignificant; for LTD significant years are: 2000-2002, 2006, 2010, 2012-2015; for NIM: 2001-2002, 2006, 2015; for PATM are: 2004-2005, 2007-2008 and 2000 for Monte Carlo only; ROA: 2002-2003, 2007; for ROE: 2001-2002, 2004, 2006-2012, 2014-2015; for RPF are: 2006-2008, and 2002 for Asump. Sig. only; significant years for SGI are: 2002, 2008, 2010, and 2001, 2007, and 2009 for Monte Carlo Sig. only.

The *EM2* model in *Table 2*, for the same tested ratios, shows almost identical statistically significant results as for the *EM1* model in *Table 1*. This would suggest, that managers in order to outplay the strict financial rules, engage in EM by manipulating different areas of financial statements with different techniques, see e.g., Beretka (2016, 22-24), and at their discretion, in years / periods of their choosing. The results suggest that the hypotheses $H0_{(a)}$: and $H0_{(b)}$: ‘may’ be rejected for the statistically significant years. Only GMI, LTA and NIM ratios show overall statistically insignificant results for each financial year for both *EM1* and *EM2* models; therefore, the hypothesis $H0_{(b)}$: holds for GMI, LTA and NIM ratios. Statistical outputs for the 14 ratios for the *EM2* model for both Confidence Interval and Significance levels are very similar to *EM1* model. Additional test statistics on annual basis for 95% Confidence Interval and at $p = 0.05$ Significance levels for the *EM2* model are available from the author.

The 2008 Financial Crisis

This paper additionally investigates evidence of EM ‘Prior to’ and ‘After’ the 2008 financial crisis. Hypothesis $H0_{(b)}$: was tested with the K-S test statistics by applying the *EM1* and the *EM2* models with the 14 bank specific ratios. In order to analyse ‘Prior to’ and ‘After’ 2008 financial crisis, ‘Prior to’ 2008 years were set from 2001 to 2007, whereas ‘After’ the financial crisis years are from 2008 to 2014. As it can be seen in *Table 1*, only Gross Margin Index (GMI), Loans to Total Assets (LTA) at $p = 0.01$ Significance Levels have statistically insignificant outputs. Same is true for tests run with the *EM2* model, statistically insignificant outputs at $p = 0.01$ are for GMI and LTA ratio and at $p = 0.05$ LTA ratio only. Statistical test outputs for *EM1* and *EM2* models for significance levels of $p = 0.05$ are available from the author. The insignificance of the outputs for GMI ratio suggests that it is less likely that it is being manipulated. The statistical insignificance of LTA suggests that there are no outstanding loans and that the liquidity of the firms is well managed. This was not the case, as the LTA statistical outputs do not show the reality of the credit institutions’ operations, which had poor liquidity and high lending, as the DTE and ETL statistical results point out. Evidence bearing out that the LTA statistical outputs are false is shown in the lending practices to Households, Consumer and Corporate Sectors in Figures 3.1, 3.1, 3.3, Beretka (2016, 106-110). Lending was in significant increase from 2003 within all three sectors with its peak in 2009 that prompted illiquidity of the lending institutions and forced bailouts of some credit institutions trading in Hungary. From 2009, astonishingly, lending maintained its level for all three sectors. In summary, it can be argued, given by the evidence from *Tables 1 and 2*, that managers engaged in EM ‘Prior to’ the 2008 financial crisis possibly to maintain parent companies’ profit expectations, thus kept engaging in EM even ‘After’ the 2008 financial crisis. Tests statistics show clear evidence that hypothesis $H0_{(b)}$: ‘may’ be rejected.

Benchmark Analysis

Earlier evidence suggests that research papers seldom test Benchmark when investigating EM. *Tables 3 and 4* present Descriptive Statistics output for Benchmark and for Base ratios. Only a handful of papers argue, see e.g., Dechow, Ge and Schrand (2010, 351), in favour of including Benchmark analysis with test statistics and/or with a combination of histograms, see e.g., Beretka (2016, 189-199). By comparing each ratio between the Base and Benchmark outputs in *Tables 3 and 4*, it can be concluded that results differ, thus indicate that there is a significant difference for Mean and S.D. are for DTE and for LTD. Observing only S.D. for ETL, PATM and ROE we may conclude that outputs show high values, suggesting that they are widely spread and are less reliable; whereas high Means for DTE, LTD suggest that they are spread away from the central point. Skewness and Kurtosis are positive and negative for both Benchmark and Base ratios outputs. Negative Skewness for ROA and ROE in the Benchmark output suggest that the values are on the right of the distribution, whereas positive values are on the left. Positive Kurtosis for benchmark output in *Table 3* are positive thus suggesting ‘pony and heavily-tailed distribution,

Table 3. Benchmark Descriptive Statistics

	N	Min	Max	Mean	Mean Std. Error	S.D.	S.D. Variance	Skew- ness	Skew Std. Error	Kurt- osis	Kurt Std. Err.
DTE	17	419.4	6023	1608	308.4	1271	1617539	2.87	0.55	9.78	1.06
ETL	16	0.02	9.23	0.68	0.57	2.28	5.21	3.99	0.56	15.9	1.09
LTD	17	1.41	1807	228	107	444	197111	3.18	0.55	11.1	1.06
LTA	17	0.18	1.85	0.59	0.08	0.34	0.12	3.45	0.55	13.7	1.06
GYEA	17	0.07	0.96	0.49	0.06	0.27	0.07	-0.4	0.55	-0.67	1.06
RPF	17	0.00	0.53	0.06	0.03	0.12	0.02	3.85	0.55	15.38	1.06
SGI	16	0.87	12.5	2.01	0.71	2.84	8.07	3.82	0.56	14.93	1.09
IRI	16	0.19	3.04	1.18	0.14	0.57	0.33	2.18	0.56	8.40	1.09
GMI	16	0.78	1.25	0.98	0.03	0.12	0.01	0.39	0.56	0.10	1.09
NIM	16	0.05	0.40	0.29	0.03	0.12	0.01	-1.3	0.56	0.51	1.09
PATM	17	-3.54	8.87	0.15	0.60	2.47	6.08	2.85	0.55	11.13	1.06
ROE	16	-20.1	23.5	-0.5	2.99	11.9	142.9	0.21	0.56	-0.08	1.09
ROA	16	-0.02	0.02	0.00	0.00	0.01	0.00	-0.3	0.56	-0.33	1.09
ETA	16	0.01	0.22	0.03	0.01	0.05	0.00	3.93	0.56	15.60	1.09

Notes: Tests ran for the period 1999 - 2015.

while negative values suggest ‘flat’ distribution, see e.g., Field (2009, 138). However, Benchmark results should be read with caution, due to the specifics of the Hungarian Accounting Standards and the lack of quality research availability. *Table 3* Benchmark outputs were calculated from the same sample as the Base ratios statistics. Despite the lack of research material, Benchmark comparisons should not be excluded from research.

Table 4. Descriptive Statistics

	N	Min	Max	Mean	Mean Std. Error	S.D.	S.D. Var.	Skew- ness	Skew Std. Err.	Kur- tosis	Kur Std. Err.
DTE	568	0.00	152260	1566	315	7524	56610025	15.4	0.10	290	0.20
ETL	515	0.00	349.6	0.78	0.68	15.42	237.66	22.6	0.11	513	0.21
LTD	579	0.00	72149	257	134	3225	10403539	19.8	0.10	432	0.20
LTA	669	0.00	53.06	0.60	0.08	2.05	4.20	25.1	0.09	644	0.19
GYEA	682	0.01	13.63	0.49	0.02	0.65	0.42	12.9	0.09	250	0.19
RPF	664	-0.0	19.12	0.07	0.03	0.74	0.55	25.3	0.09	650	0.19
SGI	592	0.00	25.56	1.89	0.12	2.97	8.85	4.5	0.10	22.8	0.20
IRI	584	0.00	84.31	1.24	0.15	3.58	12.85	21.5	0.10	497	0.20
GMI	588	-2.3	5.38	0.99	0.02	0.48	0.23	1.0	0.10	29.6	0.20
NIM	591	-0.0	3.30	0.29	0.01	0.31	0.09	3.5	0.10	20.2	0.20
PATM	673	-118	324.4	0.06	0.52	13.47	181.4	19.8	0.09	511	0.19
ROE	502	-2062	1098	-0.4	5.46	122.4	14984	-8.9	0.11	179	0.22
ROA	596	-0.3	0.18	0.00	0.00	0.04	0.00	-2.4	0.10	16.6	0.20
ETA	502	0.00	5.89	0.03	0.01	0.26	0.07	21.8	0.11	486	0.22

Notes: Descriptive Statistics for Base Ratios for the period 1999 - 2015. No modelling was performed.

Discussion

The Distribution of Ratios research approach applies a technique to use all possible computable variables from the financial statements, thus ratios that do not contain assets and/or liabilities, as component in their formula, it eliminates the reversal accrual impact and increases the credibility and power of the statistical test results in comparison to accrual testing models. By rejecting Hypotheses $H0_{(a)}$ and $H0_{(b)}$ for ratios that comprise current assets or total assets (i.e. ETA, LTA, ROA), or current / long term liabilities (i.e. DTE, ETL, LTA, LTD), which may have influenced the test statistics due to reversal accruals, a component of assets and/or liabilities in the balance sheet, this study could have made a *Type I* error.

Conclusion

This paper examined an alternative research approach, the Distribution of the Ratios method in search of evidence of Earnings Management of annual accounts of the Hungarian credit institutions (banks). Test outputs run on an annual basis with the EM1 and the EM2 models for each ratio present more details, specifically highlighting that evidence of Earnings Management, thus rejecting $H0_{(a)}$. Furthermore, evidence for EM1 and EM2 models present an opportunistic approach by managers who manipulate accounting entries at their discretion in different time frames and in all areas of the financial statements, thus $H0_{(b)}$ is rejected. Benchmark analysis was performed with the same sample as the base ratios data. Due to lack of benchmark material for financials and for credit institutions in Hungary, the results should be read with caution. However, when industry benchmark data is available, it should be used for comparison. This paper also investigated the 2008 financial crisis, and concluded that ‘Prior to’ and ‘After’ the 2008 financial crisis there is evidence of Earnings Management, thus $H0_{(b)}$ is rejected. The *EM1* and *EM2* models test results on annual basis show statistical significant results for 99% confidence and $p = 0.01$ significant levels for all 14 tested ratios. Test results for 14 ratios for the *EM1* and *EM2* models tested with all sample and with annual data of 95% confidence interval and $p = 0.05$ significance level are statistically significant and they are available from the author.

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